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Meta-analysis of transanal versus laparoscopic total mesorectal excision for rectal cancer: a 'New Health Technology' assessment in South Korea

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Purpose: Under the South Korea's unique health insurance structure, any new surgical technology must be evaluated first by the government in order to consider whether that particular technology can be applied to patients for further clinical trials as categorized as 'New Health Technology,' then potentially covered by the insurance sometime later. The aim of this meta-analysis was to assess the safety and efficacy of transanal total mesorectal excision (TaTME) for rectal cancer, activated by the National Evidence-based Healthcare Collaborating Agency (NECA) TaTME committee.

Methods: We systematically searched Ovid-MEDLINE, Ovid-Embase, Cochrane, and Korean databases (from their inception until August 31, 2019) for studies published that compare TaTME with laparoscopic total mesorectal excision (LaTME). End-points included perioperative and pathological outcomes.

Results: Sixteen cohort studies (7 for case-matched studies) were identified, comprising 1,923 patients (938 TaTMEs and 985 LaTMEs). Regarding perioperative outcomes, the conversion rate was significantly lower in TaTME (risk ratio, 0.19; 95% confidence interval, 0.11–0.34; P < 0.001); whereas other perioperative outcomes were similar to LaTME. There were no statistically significant differences in pathological results between the 2 procedures.

Conclusion: Our meta-analysis showed comparable results in preoperative and pathologic outcomes between TaTME and LaTME, and indicated the benefit of TaTME with low conversion. Extensive evaluations of well-designed, multicenter randomized controlled trials are required to come to unequivocal conclusions, but the results showed that TaTME is a potentially beneficial technique in some specific cases. This meta-analysis suggests that TaTME can be performed for rectal cancer patients as a 'New Health Technology' endorsed by NECA in South Korea. **[Ann Surg Treat Res 2021;101(3):167-180]**

Key Words: Colorectal neoplasms, Laparoscopic total mesorectal excision, Meta-analysis, Systematic review, Transanal total mesorectal excision

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INTRODUCTION

Colorectal cancer is the third most common diagnosis and the third-leading cause of death worldwide [1]. Total mesorectal excision (TME) was first described by Heald and Ryall in 1982 [2], and it has become the standard treatment for rectal cancer by effectively reducing recurrence, improving quality of life, and prolonging survival [3]. Laparoscopic surgery has replaced open surgery, allowing favorable short-term outcomes, particularly for pain reduction, reduced blood loss, and rapid recovery [4]. Laparoscopic surgery has been widely accepted and used for minimally invasive surgical procedures; however, conversions from laparoscopic TME (LaTME) to open procedure have been common in patients with rectal cancer [5]. The main reasons for the unsatisfactory conversion rate were the patient characteristics of male sex (because of the narrow pelvis) and the high body mass index (BMI) [6,7]. Additionally, high incidence of positive resection margin and poor quality of TME specimen, which are factors known as poor oncologic outcomes in deep narrow pelvis, exist [8].

To address the aforementioned limitations, transanal TME (TaTME) was first described by Sylla et al. [9]. In several studies, this "bottom-up" approach for mid-low rectal cancer has been proposed as a safe and effective technique for patients by demonstrating the benefits of more precise excision under appropriate visualization and potentially improving specimen quality and resection margins [10-12]. However, it still remains controversial because its oncological feasibility and safety have not yet been validated by large randomized controlled trials (RCTs). Recent studies reported that the TaTME approach showed unfavorable local recurrence rates and a higher risk of anastomotic leak than a nontransanal approach [13,14]. TaTME also seems to be associated with substantial morbidities such as urethral and other urologic injuries [15]. It is suggested that these complications are related to the surgeon's lack of surgical training or learning curves [16,17].

The national health insurance system of South Korea is unique public health insurance where all citizens are forced to sign in, and all insurance policyholders are burdened with the duty to pay for the insurance. By the unique structure of South Korea's health insurance system, if a new medical procedure doesn't pass the New Health Technology Assessment (nHTA), it cannot be covered by the health insurance, thus not being available to policyholders even as an uncovered procedure. Therefore, nHTA is an essential step undertaken to evaluate whether a procedure's clinical safety/efficacy is appropriate to be used on insurance policyholders through a systematic review.

TaTME is a new surgical technique in which the technical method of conventional LaTME has been changed and is subject to nHTA. Thus, a systematic review and meta-analysis of the current data from the latest and most convincing studies was conducted, comparing the safety and efficacy between TaTME and LaTME for nHTA by the National Evidence-based Healthcare Collaborating Agency (NECA) committee.

METHODS

Registration and search strategy

This study was conducted in accordance with the PRISMA (preferred reporting items for systematic reviews and metaanalyses) guidelines [18]. The protocol was registered in PROSPERO under the number CRD42021230076. Comprehensive searches were performed through the databases of MEDLINE, Embase, Cochrane Library, and Korean databases (KoreaMed, KMbase, KISTI, KISS, and RISS) from their inception until August 31, 2019. The MEDLINE and Embase databases were searched using the following terms with Boolean operators: (rectal neoplasms OR rectal tumor OR rectal cancer OR colorectal neoplasms OR colorectal tumor OR colorectal cancer) AND (transanal total mesorectal excision OR TaTME OR Ta-TME OR transanal minimally invasive surgery OR TAMIS OR transanal endoscopic surgery OR transanal rectal resection). The Cochrane database was searched using the following keyword: transanal total mesorectal excision. The Korean database was searched using the following terms with Boolean operators: neoplasms AND transanal total mesorectal excision.

Inclusion and exclusion criteria

In accordance with PICOS (patient, intervention, comparator, outcome, and study design) criteria, studies were included if they met the following inclusion criteria: (1) patients diagnosed with rectal cancer; (2) patients underwent either TaTME or LaTME: (3) at least 1 of the following perioperative outcomes or pathological data were available; and (4) studies were cohort studies. The exclusion criteria were as follows: (1) reviews, letters, editorials, commentaries, conference abstract, and clinical reports; (2) studies with sample size below 20 in each group: (3) languages other than English; (4) inappropriate data; (5) duplicate patient series; (6) inadequate technique for intervention or comparator; and (7) nonhuman researches.

Data extraction and quality assessment

The following data were extracted by the 2 reviewers (SHK and YIJ). (1) Demographic data: study design, sex, BMI, age, American Society of Anesthesiologists physical status classification, tumor location, and neoadjuvant treatment; (2) Perioperative outcomes: operation time, intraoperative blood loss, conversion rate, hospital stay, readmission, reoperation, diverting ileostomy, major complication (Clavien-Dindo classification III–V), anastomotic leakage, intestinal

Study	Year	Country	Case- matched	Sex, mali	e:female	Body mass ii	ndex (kg/m ²)	Age	(yr)	ASA 5 I + II/II	core, I + IV	Tumor	Neoad therapy,	juvant yes/no	Therapy type
			study	TaTME	LaTME	TaTME	LaTME	TaTME	LaTME	TaTME	LaTME	וחרמווחוו	TaTME	LaTME	(duration)
Fernández-Hevia et al. [24]	a 2015	Spain	No	24:13	22:15	23.7 ± 3.6 (18.0-31.0)	25.1 ± 4.0 (15.0-31.6)	64.5 ± 11.8	69.5 ± 10.5	30/7	25/12	Low/mid	28/9	23/14	CRT (NR)
Chen et al. [25]	2016	Taiwan	Yes	38:12	76:24	24.2 ± 3.7 (16.0-37.0)	24.6 ± 3.1 (17.2-32.8)	57.3 ± 11.9 (29-80)	58.3 ± 11.3 (34-82)	33/17	69/31	Low/mid	50/0	100/0	CRT (NR)
Rasulov et al. [26]	2016	Russia	No	11:11	14:9	26.0 (19.7–32.3)	26.0 (18.3–37.2)	56.0 (30.0–69.0)	60.0 (15.0–78.0)	NR	NR	Low/mid	19/3	11/12	CRT (SC/LC)
Lelong et al. [27] 2017	France	No	23:11	22:16	24.0 (18.6–45.0)	24.2 (17.7–32.7)	NR	NR	30/4	36/2	Low	30/4	35/3	CRT (NR)
Chang and Kiu [28]	2018	Taiwan	Yes	13:10	13:10	25.8 ± 4.3	25.0 ± 3.9	62.4 ± 12.9	62.9 ± 12.6	20/3	22/1	Low	8/15	14/9	CRT (NR)
Detering et al. [29]	2018	Netherlands	Yes	288:108	281:115	NR	NR	NR	NR	330/66	331/65	Low/mid	255/140	252/143	CRT (SC/LC)
Mege et al. [30]	2018	France	Yes	23:11	23:11	25.0 ± 4.0	25.0 ± 3.0	58.0 ± 14.0	59.0 ± 13.0	32/1	32/2	Low	29/5	29/5	CRT (LC)
Persiani et al. [31]	2018	Italy	Yes	30:16	31:15	25.0 (19.1–32.8)	25.6 (18.8–33.4)	69.0 (36.0–94.0)	66.5 (28.0–86.0)	NR	NR	Low/mid	26/20	32/14	CRT (NR)
Alamili et al. [32]	2019	Denmark	No	26:14	8:12	26.0 (18.0–38.0)	24.0 (19.0–29.0)	69.0 (48.0–89.0)	73.0 (50.0–84.0)	34/6	17/3	Low/mid/ high	13/27	4/16	CRT (NR)
Bjoern et al. [33.] 2019	Denmark	No	37:12	16:20	26.6 ± 3.5	25.5 ± 4.8	64.9 ± 9.6	62.4 ± 10.1	41/8	35/1	Low/mid	8/41	8/28	CRT (NR)
Chen et al. [34]	2019	Taiwan	No	29:10	42:22	25.4 ± 4.0	24.5 ± 3.3	62.0 ± 14.9	64.0 ± 12.2	33/6	58/6	Low	15/24	31/33	CRT (NR)
Perdawood et al [35]	. 2019	Denmark	No	22:7	17:12	26.8 ± 4.5	26.3 ± 4.8	70.0 ± 7.1	70.1 ± 8.4	20/9	27/2	Low/mid	5/24	4/25	CRT (NR)
Roodbeen et al. [36]	2019	Netherlands	Yes	34:7	32:9	26.7 ± 1.9 (20.9-32.3)	26.1 ± 4.0 (19.4-36.0)	62.5 ± 10.7 (33-87)	66.0 ± 9.2 (48-83)	36/5	38/3	Low	18/23	18/23	RT, CT, CRT (NR)
Rubinkiewicz et al. [37]	2019	Poland	No	13:10	13:10	26.0 (22.8–29.7)	26.5 (23.8–30.6)	60.0 (51.0–67.0)	64.0 (58.0–67.0)	18/5	18/5	Low	13/10	15/8	CRT (LC)
Sparreboom et al. [38]	2019	Belgium Netherlands	Yes	33:15	32:16	27.0 (24.5–30.7)	26.1 (24.0–29.0)	65.0 (56.8–71.0)	64.0 (59.3–73.0)	33/15	34/14	Low/mid/ high	NR	NR	RT, CT (SC/LC)
Veltcamp et al. [39]	2019	Netherlands	No	18:9	20:7	27.6 (25.7–29.5)	26.1 (25.1–27.3)	68.0 (64.4–71.6)	62.7 (59.6–65.7)	25/2	25/2	Low/mid/ high	22/5	18/7	RT, CRT (NR)
Values are prese ASA, American 5 SC, short-course;	nted as society : LC, Ioı	number only, i of anesthesiold ng-course; RT, u	mean ± st ogists; Ta1 radiothera	tandard de IME, trans apy; CT, cl	eviation (re sanal total hemothere	ange), or medi mesorectal e: apy.	an (range). xcision; LaTMI	E, laparoscopic	c total mesored	tal excis	ion; CR1	, chemorad	iotherapy	/; NR, no	t reported;

Table 1. Demographics and clinical characteristics of the included studies

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obstruction, ureter or urethral injury, urinary retention, urinary tract infection, and mortality: (3) Pathological outcomes: circumferential resection margin (CRM) and distal resection margin (DRM) involvement, length of CRM and DRM, incompleteness of mesorectum, and harvested lymph nodes. The methodological quality and risk of bias of the included studies were assessed using the Newcastle–Ottawa quality assessment scale [19]. Any disagreements were settled by consensus-based discussion between the 2 reviewers (SHK and YIJ).

Definitions

Inappropriate techniques described in the exclusion criteria are defined as all procedures except interventional (TaTME) or comparative procedures (LaTME). The location of the tumor is categorized as low (0–5 cm from the anal verge), middle (5.1–10 cm from the anal verge), and high (10.1–15 cm from the anal verge). Conversion in LaTME was defined when the procedure was completed with open surgery. Conversion in TaTME was defined as a case in which the procedure was completed by open surgery, or the TME was performed by the transanal approach but the conversion occurred at the transabdominal phase. Mesorectal resection quality was scored using 3 grades complete, nearly complete, or incomplete, as defined by Quirke et al. [20].

Statistical analysis

Meta-analyses were performed with Review Manager 5.3 software (Cochrane Collaboration, Oxford, UK). Dichotomous data were pooled as risk ratios (RRs) with 95% confidence intervals (CIs) using the Mantel-Haenszel method, which can avoid biased estimates by incorporating evidence from single zero studies without requiring the standard continuity correction [21]. Mean differences (MDs) and 95% CI were pooled for continuous variables using the inverse variance method. If the median and range were reported instead of the mean and standard deviation, these values were estimated using the method devised by Hozo et al. [22]. The Q test and I^2 statistic were used to evaluate heterogeneity among studies. The Cochrane Handbook for Systematic Reviews of Interventions provides a rule of thumb for interpreting the I^2 statistic; that is $I^2 \le 40\%$ may indicate unimportant heterogeneity, $30\% \le I^2 \le$ 60% may represent moderate heterogeneity, 50% \leq I^2 \leq 90% may represent substantial heterogeneity, and $70\% \le I^2 \le 100\%$ implies that heterogeneity may be considerable. A Cochrane Q statistical P < 0.100 and/or I^2 > 50% was taken to indicate significant heterogeneity, in which case a random-effects model was applied [23]. Otherwise, a fixed-effects model was employed. Funnel plots were used in meta-analysis to visually detect the presence of publication bias.

RESULTS

Study characteristics

As a result of the literature search, 1,940 studies were identified. Among the search results obtained through the online databases in the first screening, 534 duplicate articles were removed. After that, through the review of the abstracts, 857 articles that did not fit the language and article type were excluded. The remaining articles were fully checked to ensure they met the inclusion criteria, and 533 additional articles were excluded. Finally, 16 cohort studies met all the inclusion criteria and none of the exclusion criteria (Fig. 1) [24-39]. Seven of the cohort studies were case-matched studies comparing TaTME with LaTME for rectal cancer [25,28-31,36,38]. The patient demographics, clinical characteristics, and quality assessment scores are shown in Table 1 and Supplementary Table 1.

Perioperative outcomes

A total of 12 studies [24-28,30-32,34-36,38] reported operation time in 946 patients, and we found no statistically significant difference between TaTME and LaTME groups (MD, 7.52; 95% CI, -10.03-25.07; P = 0.400) (Fig. 2A). Perdawood et al. [35] and Fernández-Hevia et al. [24] reported a 2-team approach as standard; Alamili et al. [32], Chen et al. [34], Roodbeen et al. [36], Sparreboom et al. [38], Chang and Kiu [28], and Mege et al. [30] used a 1-team approach. Persiani et al. [31], Lelong et al. [37], Chen et al. [25], and Rasulov et al. [26] used both a 1- and 2-team approach. There was a notable heterogeneity between included studies ($I^2 = 88\%$, P < 0.001).

Intraoperative blood loss reported in 6 studies [25.26.28.32.34.35] that investigated 462 patients (Fig. 2B). There was no statistically significant difference between groups (MD, -8.37; 95% CI, -27.82-11.08; P = 0.410), and moderate heterogeneity was existed (I² = 48%, P = 0.090).

Data on conversion rate were extracted from 13 studies [24-32,34-36,38] assessing 1,738 patients (Fig. 2C). Upon analysis, the TaTME group displayed significant lower conversion rate compared to the LaTME group (RR, 0.19; 95% CI, 0.11–0.34; P < 0.001). There was no heterogeneity between included studies (I² = 0%, P = 0.460).

Twelve studies [24-32,34,36,38] assessed hospital stay in 1,680 patients (Fig. 3A). Pooled analysis indicated that the result was not statistically significant between the 2 groups (MD, -0.51; 95% CI, -1.43-0.41; P = 0.280). Substantial heterogeneity existed in hospital stay (I² = 88%, P < 0.001).

A total of 7 studies [24,25,27,29,32,36,38] reported data on readmission in 1,326 patients (Fig. 3B). There was no statistically significant difference between the 2 groups (RR, 0.90; 95% CI, 0.70–1.15; P = 0.400), and moderate heterogeneity existed ($I^2 = 49\%$, P = 0.070).

Five studies [24,25,27,32,38] that evaluated 452 patients were



Fig. 1. PRISMA diagram of the search strategy. TaTME, transanal total mesorectal excision; LaTME, laparoscopic total mesorectal excision.

pooled for analysis of reoperation (Fig. 3C). The difference was not statistically significant between the 2 groups (RR, 0.86; 95% CI, 0.45–1.61; P = 0.630). There was no heterogeneity between included studies ($I^2 = 0\%$, P = 0.470).

Analysis of the major complication (Clavien-Dindo classification III–V) was performed based on 9 studies [24,26,27,30,31,36-39] evaluating 629 patients (Fig. 4A). There was no statistically detected difference between the groups (RR, 0.75; 95% CI, 0.50–1.13; P = 0.170). Heterogeneity in this regard was not significant ($I^2 = 0\%$, P = 0.640).

Data on anastomotic leakage were available in 9 studies [24,28-32,34,36,38], which assessed 1,196 patients (Fig. 4B). We found that there was no statistically significant difference between the 2 groups (RR, 1.32; 95% CI, 0.96–1.82; P = 0.090), with no heterogeneity ($I^2 = 26\%$, P = 0.210).

Intestinal obstruction was investigated in 8 studies [24,25,28,30-32,34,38], and remained unchanged between the 2 groups (RR, 1.01; 95% CI, 0.59–1.72; P = 0.970) (Fig. 4C). There was no heterogeneity between included studies ($I^2 = 0\%$, P = 0.940).

Moreover, the occurrence of other complications such as diverting ileostomy (RR, 1.20; 95% CI, 0.87–1.65; P = 0.270; I² = 85%), mortality within 30 days (RR, 0.40; 95% CI, 0.06–2.73; P = 0.350; I² = 0%), ureter or urethral injury (RR, 0.83; 95% CI, 0.10–6.63; P = 0.860; I² = 0%), urinary retention (RR, 1.23; 95% CI, 0.04–38.46; P = 0.910; I² = 74%), and urinary tract infection (RR, 1.08; 95% CI, 0.29–4.11; P = 0.910; I² = 0%) were similar

between the 2 groups (Supplementary Fig. 1, 2).

Pathological outcomes

Fifteen studies [24-31,33-39] that assessed 1,863 patients reported on CRM involvement (Fig. 5A). The results showed that CRM involvement was similar between the 2 groups (RR, 0.73; 95% CI, 0.48–1.10; P = 0.130), and no heterogeneity was detected between included studies ($I^2 = 0\%$; P = 0.750).

The length of CRM was reported in 4 studies [24,32,35,36] that investigated 274 patients (Fig. 5B). The comparison between the 2 groups resulted in a difference that was not statistically significant (MD, 1.64; 95% CI, -1.32-4.60; P = 0.280), but there was a notable heterogeneity ($I^2 = 97\%$; P < 0.001).

Seven studies [26,27,30,35-38] assessed 467 patients and reported on DRM involvement (Fig. 5C). DRM involvement is not significantly different between the 2 groups (RR, 0.60; 95% CI, 0.32–1.15; P = 0.120), with no significant between-study heterogeneity ($I^2 = 0\%$, P = 0.880).

A total of 9 studies [24.25.28.30-32.34-36] reported length of DRM in 733 patients (Fig. 5D), which was not significantly different between the 2 groups (MD, 1.88; 95% CI, -2.96-6.73; P = 0.450). Statistical heterogeneity among the studies was high ($I^2 = 87\%$, P < 0.001).

Eleven studies [24,26,27,30-33,35-37,39] assessed 736 patients and provided data on incompleteness of mesorectum (Fig. 5E). Pooled analysis showed that the incompleteness of mesorectum was equivalent between the 2 groups (RR, 1.00: 95% CI,

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Α	-	TaTME			этмғ			Mean difference		Meand	lifforor	100	
Study or subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, random, 95% Cl		IV, rando	om, 95	% CI	
Alamili 2019	349	76.75	40	268.5	74	20	6.8%	80.50 [40.28, 120.72]					
Chang 2018	200	57.4	23	191.8	64.8	23	7.4%	8.20 [-27.18, 43.58]			-		
Chen 2016	182.1	55.4	50	178.7	34.8	100	9.7%	3.40 [-13.40, 20.20]		_			
Chen 2019	210	57	39	184	55	64	9.1%	26.00 [3.60, 48.40]					
Fernández-Hevia 2015	215	60	37	252	50	37	8.7%	-37.00 [-62.17, -11.83]					
Lelong 2017	532	97.5	34	576	82.5	38	6.6%	-44.00 [-85.98, -2.02]			-		
Mege 2018	246	48	34	247	60	34	8.6%	-1.00 [-26.83, 24.83]					
Perdawood 2019	296.97	72.64	29	355.83	92.92	29	6.5%	-58.86 [-101.79, -15.93]					
Persiani 2018	276	65	46	272	68	46	8.5%	4.00 [-23.18, 31.18]				_	
Rasulov 2016	350	57.5	22	318.75	61.25	23	7.5%	31.25 [-3.45, 65.95]					
Roodbeen 2019	318	26.25	41	300	34.5	41	10.0%	18.00 [4.73, 31.27]				_	
Sparreboom 2019	221	19	48	180	16	48	10.5%	41.00 [33.97, 48.03]					
Total (95% CI)			443			503	100.0%	7.52 [-10.03, 25.07]					
Heterogeneity: $Tau^2 = 75$	53.83; Cł	ni ² = 90	.17, d	f = 11 (F	o < 0.0	01); I ²	= 88%		H				
Test for overall effect: Z	= 0.84 (F	P = 0.40	D)						-100	-50	0	50	100
									Favo	JURS TATIVIE	ra	vours La	INE

В	т	aTMF			aTMF			Mean differen	ICA	Mean difference
Study or subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, fixed, 95%	CI	IV, fixed, 95% CI
Alamili 2019	100	800	40	212	162.5	20	0.6%	-112.00 [-369.94, 145.9	94]	←
Chang 2018	39.1	63.9	23	36.9	77.2	23	22.6%	2.20 [-38.76, 43.4	16]	
Chen 2016	68.9	89.6	50	88.2	102.5	100	37.1%	-19.30 [-51.24, 12.6	64]	
Chen 2019	63	102	39	42	59	64	30.7%	21.00 [-14.12, 56.4	12]	+
Perdawood 2019	73.79	56.23	29	188.97	283.63	29	3.4%	-115.18 [-220.42, -9.9	94] ·	←
Rasulov 2016	135	105	22	197.5	167.5	23	5.7%	-62.50 [-143.81, 18.8	81]	
Total (95% CI)			203			259	100.0%	-8.37 [-27.82, 11.0	08]	•
Heterogeneity: Chi ² = 9.	67, df = 5	5 (P = 0).09);	$l^2 = 489$	%				ł	
Test for overall effect: Z	= 0.84 (F	9 = 0.40	D)							
										Favours la IME Favours La IME

С	тати		LaT			Pick ratio	Pick	atio
Study or subgroup	Events	Total	Events	Total	Weight	M-H, fixed, 95% Cl	M-H, fixed	I, 95% CI
Alamili 2019	1	40	1	20	1.7%	0.50 [0.03, 7.59]		
Chang 2018	0	23	0	23		Not estimable		
Chen 2016	1	50	5	100	4.4%	0.40 [0.05, 3.33]		
Chen 2019	1	39	1	64	1.0%	1.64 [0.11, 25.49]		
Detering 2018	6	396	34	396	44.5%	0.18 [0.07, 0.42]		
Fernández-Hevia 2015	0	37	0	37		Not estimable		
Lelong 2017	1	34	9	38	11.1%	0.12 [0.02, 0.93]		
Mege 2018	1	34	0	34	0.7%	3.00 [0.13, 71.15]		
Perdawood 2019	0	29	2	29	3.3%	0.20 [0.01, 3.99]		
Persiani 2018	0	46	9	46	12.4%	0.05 [0.00, 0.88]	←	
Rasulov 2016	1	22	1	23	1.3%	1.05 [0.07, 15.70]		
Roodbeen 2019	0	41	9	41	12.4%	0.05 [0.00, 0.88]	←	
Sparreboom 2019	0	48	5	48	7.2%	0.09 [0.01, 1.60]	•	_
Total (95% CI)		839		899	100.0%	0.19 [0.11, 0.34]	•	
Total events	12		76					
Heterogeneity: Chi ² = 9.7	'6, df = 10 (F	P = 0.46);	$l^2 = 0\%$				0.01 0.1 1	10 100
Test for overall effect: Z =	= 5.84 (P < 0	.001)					Favours TaTME	Favours LaTME

Fig. 2. Forest plots of risk ratios and mean differences of perioperative outcomes. (A) Operative time, (B) intraoperative blood loss and, (C) conversion rate. A random-effect model was used for meta-analysis of operative time. Fixed-effects models were used for meta-analysis of intraoperative blood loss and conversion rate. Mean differences and risk ratios are shown with 95% confidence intervals (CIs). TaTME, transanal total mesorectal excision; LaTME, laparoscopic total mesorectal excision; SD, standard deviation; IV, inverse variance; df, degree of freedom.

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~	Т	aTME		L	.aTME			Mean difference		Mean	differend	се	
Study or subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, random, 95% CI		IV, rand	dom, 95%	S CI	
Alamili 2019	13	6.5	40	12.5	6.5	20	4.4%	0.50 [-2.99, 3.99]	_				
Chang 2018	9.7	3.2	23	9.4	3.6	23	7.7%	0.30 [-1.67, 2.27]					
Chen 2016	7.4	2.5	50	7.1	3.8	100	10.4%	0.30 [-0.72, 1.32]		-		-	
Chen 2019	9.2	2.7	39	9.6	4.6	64	9.4%	-0.40 [-1.81, 1.01]					
Detering 2018	6.5	1.3	396	6	0.8	396	11.9%	0.50 [0.35, 0.65]			-0-		
Fernández-Hevia 2015	6.8	3	37	9	7.6	37	6.1%	-2.20 [-4.83, 0.43]	•		<u> </u>		
Lelong 2017	8	4.25	34	9	4.5	38	7.6%	-1.00 [-3.02, 1.02]	_				
Mege 2018	10	6	34	11	5	34	6.1%	-1.00 [-3.63, 1.02]			_	_	
Persiani 2018	5	3.25	46	7	8.25	46	6.2%	-2.00 [-4.56, 0.56]			<u> </u>		
Rasulov 2016	10	3	22	9.25	3.25	23	8.1%	0.75 [-1.08, 2.58]					
Roodbeen 2019	8	1	41	11	2.25	41	11.1%	-3.00 [-3.75, -2.25]		_			
Sparreboom 2019	8	1.88	48	7.5	2.2	48	10.9%	0.50 [-0.32, 1.32]				-	
Total (95% CI)			810			870	100.0%	-0.51 [-1.43, 0.41]					
Heterogeneity: Tau ² = 1.8	35; Chi ² :	= 90.68	8, df =	11 (P <	0.001); ² = 8	88%						
Test for overall effect: Z =	= 1.09 (P	9 = 0.28	3)						-4		U		4
									Favo	ours la IME	Fave	ours La H	VIE

В	TaTI	ME	l aTI	ME		Risk ratio		Risk	ratio	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, fixed, 95% Cl		M-H, fixe	d, 95% Cl	
Alamili 2019	11	40	1	20	1.2%	5.50 [0.76, 39.66]		_		
Chen 2016	3	50	10	100	6.1%	0.60 [0.17, 2.08]				
Detering 2018	68	396	74	396	67.8%	0.92 [0.68, 1.24]			ŀ	
Fernández-Hevia 2015	2	37	8	37	7.3%	0.25 [0.06, 1.10]				
Lelong 2017	0	34	6	38	5.6%	0.09 [0.01, 1.47]	←	-	_	
Roodbeen 2019	6	41	8	41	7.3%	0.75 [0.29, 1.97]				
Sparreboom 2019	10	48	5	48	4.6%	2.00 [0.74, 5.42]		-		
Total (95% CI)		646		680	100.0%	0.90 [0.70, 1.15]		•		
Total events	100		112							
Heterogeneity: Chi ² = 11.	77, df = 6 (F	P = 0.07);	$l^2 = 49\%$						10	100
Test for overall effect: Z =	= 0.85 (P = 0	0.40)					0.01 Fav	ours TaTME	Favours I	_aTME

	TaTI	ME	LaTI	ME		Risk ratio			Risk ra	tio		
Study or subgroup	Events	Total	Events	Total	Weight	M-H, fixed, 95% Cl		M-H	, fixed,	95% CI		
Alamili 2019	6	40	2	20	14.1%	1.50 [0.33, 6.77]		-				
Chen 2016	2	50	3	100	10.6%	1.33 [0.23, 7.72]			-			
Fernández-Hevia 2015	1	37	3	37	15.9%	0.33 [0.04, 3.06]						
Lelong 2017	0	34	4	38	22.5%	0.12 [0.01, 2.22]	←			_		
Sparreboom 2019	8	48	7	48	37.0%	1.14 [0.45, 2.90]						
Total (95% CI)		209		243	100.0%	0.86 [0.45, 1.61]						
Total events	17		19									
Heterogeneity: Chi ² = 3.5	7, df = 4 (P	= 0.47); I	² = 0%				0.01	0.1	1	10	1	
Test for overall effect: Z =	= 0.48 (P = 0	0.63)					Fav	ours TaTN	ΛE .	Favours I	_aTME	:

Fig. 3. Forest plots of risk ratios and mean differences of perioperative outcomes. (A) Hospital stay, (B) readmission, and (C) reoperation. A random-effect model was used for meta-analysis of hospital stay. Fixed-effects models were used for meta-analysis of readmission and reoperation. Mean differences and risk ratios are shown with 95% confidence intervals (CIs). TaTME, transanal total mesorectal excision; LaTME, laparoscopic total mesorectal excision; SD, standard deviation; IV, inverse variance; df, degree of freedom.

0.70–1.45; P = 0.980). There was no heterogeneity between included studies ($I^2 = 10\%$, P = 0.350).

Among the 9 studies [25-28,30,32,34-36] that evaluated 684

patients, the difference in harvested lymph nodes was not statistically significant (MD, 0.44; 95% CI, -0.89-1.78; P = 0.520) (Fig. 5F). No significant heterogeneity existed between included

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~	TaTI	ME	LaTI	ME		Risk ratio		Ri	sk ratio		
Study or subgroup	Events	Total	Events	Total	Weight	M-H, fixed, 95% Cl		M−H, fi	xed, 95	% CI	
Fernández-Hevia 2015	3	37	5	37	11.0%	0.60 [0.15, 2.33]			<u> </u>		
Lelong 2017	2	34	7	38	14.5%	0.32 [0.07, 1.43]			+		
Mege 2018	4	34	2	34	4.4%	2.00 [0.39, 10.20]		_			
Persiani 2018	0	46	1	46	3.3%	0.33 [0.01, 7.98]			_		
Rasulov 2016	0	22	2	23	5.4%	0.21 [0.01, 4.12]			_	_	
Roodbeen 2019	9	41	7	41	15.4%	1.29 [0.53, 3.12]		-			
Rubinkiewicz 2019	3	23	4	23	8.8%	0.75 [0.19, 2.98]					
Sparreboom 2019	9	48	10	48	22.0%	0.90 [0.40, 2.02]		_	-		
Veltcamp 2019	3	27	7	27	15.4%	0.43 [0.12, 1.49]			+		
Total (95% CI)		312		317	100.0%	0.75 [0.50, 1.13]		•			
Total events	33		45								
Heterogeneity: Chi ² = 6.0	9, df = 8 (P	= 0.64); I	² = 0%				0.01	0.1	1	10	100
Test for overall effect: Z =	= 1.38 (P = 0).17)					Fav	ours TaTME	Fa	vours La	INF



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D	TaTM	ИE	LaTI	ИE		Risk ratio	Ris	sk ratio	
Study or subgroup	Events	Total	Events	Total	Weight	M−H, fixed, 95% Cl	M−H, fi	xed, 95% Cl	
Alamili 2019	4	40	4	20	9.3%	0.50 [0.14, 1.79]			
Chang 2018	1	23	0	23	0.9%	3.00 [0.13, 70.02]			
Chen 2019	1	39	0	64	0.7%	4.88 [0.20, 116.79]			
Detering 2018	52	315	25	287	45.7%	1.90 [1.21, 2.97]			
Fernández-Hevia 2015	2	37	4	37	7.0%	0.50 [0.10, 2.56]			
Mege 2018	1	34	5	34	8.7%	0.20 [0.02, 1.62]		<u> </u>	
Persiani 20218	3	46	2	46	3.5%	1.50 [0.26, 8.56]			
Roodbeen 2019	4	27	5	28	8.6%	0.83 [0.25, 2.77]			
Sparreboom 2019	10	48	9	48	15.7%	1.11 [0.50, 2.49]	-		
Total (95% CI)		609		587	100.0%	1.32 [0.96, 1.82]		•	
Total events	78		54						
Heterogeneity: Chi ² = 10.8	36, df = 8 (F	^o = 0.21);	$l^2 = 26\%$					1 1	0 100
Test for overall effect: Z =	1.70 (P = 0	.09)					Favours TaTME	Favours	LaTME

С	TaTI	ME	LaTI	ИF		Risk ratio	Risk ratio	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, fixed, 95% Cl	M-H, fixed, 95% Cl	I
Alamili 2019	1	40	1	20	5.5%	0.50 [0.03, 7.59]		
Chang 2018	1	23	1	23	4.1%	1.00 [0.07, 15.04]		
Chen 2016	0	50	2	100	6.9%	0.40 [0.02, 8.10]		
Chen 2019	1	39	3	64	9.4%	0.55 [0.06, 5.08]		
Fernández-Hevia 2015	4	37	2	37	8.2%	2.00 [0.39, 10.26]		-
Mege 2018	4	34	2	34	8.2%	2.00 [0.39, 10.20]		_
Persiani 2018	6	46	6	46	24.7%	1.00 [0.35, 2.87]		
Sparreboom 2019	7	48	8	48	32.9%	0.88 [0.34, 2.22]		
Total (95% CI)		317		372	100.0%	1.01 [0.59, 1.72]	•	
Total events	24		25					
Heterogeneity: Chi ² = 2.3	5, df = 7 (P	= 0.94); I	² = 0%					10 100
Test for overall effect: Z =	0.04 (P = 0	.97)					Favours TaTME Favours	s LaTME

Fig. 4. Forest plots of risk ratios of perioperative outcomes. (A) Major complications (Clavien-Dindo classification III–V), (B) anastomotic leakage, and (C) intestinal obstruction. Fixed-effects models were used for meta-analysis. Risk ratios are shown with 95% confidence intervals (CIs). TaTME, transanal total mesorectal excision; LaTME, laparoscopic total mesorectal excision; df, degree of freedom.

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А	TaTI	ME	LaTI	ME		Risk ratio	Risk ratio
Study or subgroup	Events	Total	Events	Total	Weight	M-H, fixed, 95% Cl	M−H, fixed, 95% Cl
Bjoern 2019	2	49	2	36	4.6%	0.73 [0.11, 4.97]	
Chang 2018	0	23	4	23	8.9%	0.11 [0.01, 1.95]	← ← ← ← ← ← ← ← ← ← ← ← ← ← ← ← ← ← ←
Chen 2016	2	50	10	100	13.2%	0.40 [0.09, 1.76]	
Chen 2019	0	39	5	64	8.3%	0.15 [0.01, 2.60]	← → → → → → → → → → → → → → → → → → → →
Detering 2018	17	396	16	396	31.7%	1.06 [0.54, 2.07]	_ _
Fernández-Hevia 2015	0	37	0	37		Not estimable	
Lelong 2017	2	34	4	38	7.5%	0.56 [0.11, 2.86]	
Mege 2018	4	34	5	34	9.9%	0.80 [0.23, 2.73]	
Perdawood 2019	1	29	3	29	5.9%	0.33 [0.04, 3.02]	
Persiani 2018	0	46	0	46		Not estimable	
Rasulov 2016	1	22	0	23	1.0%	3.13 [0.13, 72.99]	
Roodbeen 2019	2	41	3	41	5.9%	0.67 [0.12, 3.78]	
Rubinkiewicz 2019	1	23	0	23	1.0%	3.00 [0.13, 70.02]	
Sparreboom 2019	2	48	1	48	2.0%	2.00 [0.19, 21.33]	
Veltcamp 2019	0	27	0	27		Not estimable	
Total (95% CI)		898		965	100.0%	0.73 [0.48, 1.10]	•
Total events	34		53				
Heterogeneity: Chi ² = 7.6	2, df = 11 (F	P = 0.75);	$I^2 = 0\%$				
Test for overall effect: Z =	1.50 (P = 0).13)					
	,						Favours fativite Favours Lativite

D		TaTME		1	LaTME			Mean difference	Mear	n difference	
Study or subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, random, 95% CI	IV, ran	dom, 95% Cl	
Alamili 2019	7	8	40	10	6	20	20.7%	-3.00 [-6.61, 0.61]		<u> </u>	
Fernández-Hevia 2015	12	0.9	37	11	0.6	37	29.8%	1.00 [0.65, 1.35]		-	
Perdawood 2019	10.48	8.05	29	8	6.52	29	20.2%	2.48 [-1.29, 6.25]			
Roodbeen 2019	10	1.95	41	5	1.75	41	29.3%	5.00 [4.20, 5.80]		-8-	
Total (95% CI)			147			127	100.0%	1.64 [-1.32, 4.60]			
Heterogeneity: Tau ² = 7.6	51; Chi ² :	= 86.86	6, df = 3	6 (P < 0.	001); I ²	² = 97%					
Test for overall effect: Z =	= 1.09 (P	e = 0.28	3)						-10 -5	0 5	10

Favours TaTME Favours LaTME

С	ТаТІ		I a I			Risk ratio			Riek ra	atio		
Study or subgroup	Events Total		Events	Total	Weight	M-H, fixed, 95% Cl		M-H	, 95% CI			
Lelong 2017	0	34	1	38	6.5%	0.37 [0.02, 8.82]			-			
Mege 2018	1	34	1	34	4.6%	1.00 [0.07, 15.34]					_	
Perdawood 2019	0	29	1	29	6.9%	0.33 [0.01, 7.86]						
Rasulov 2016	5	22	5	23	22.4%	1.05 [0.35, 3.12]						
Roodbeen 2019	0	41	3	41	16.0%	0.14 [0.01, 2.68]	-					
Rubinkiewicz 2019	0	23	1	23	6.9%	0.33 [0.01, 7.78]			-			
Sparreboom 2019	5	48	8	48	36.7%	0.63 [0.22, 1.77]		-		-		
Total (95% CI)		231		236	100.0%	0.60 [0.32, 1.15]						
Total events	11		20									
Heterogeneity: Chi ² = 2.	40, df = 6 (P	= 0.88); I	² = 0%					01		1	0	100
Test for overall effect: Z	= 1.55 (P = 0	0.12)					5.01 Fav	ours TaT	ME	Favours	: LaT	ME

Fig. 5. Forest plots of risk ratios and mean differences of pathological outcomes. (A) Circumferential resection margin involvement, (B) length of circumferential resection margin, (C) distal resection margin involvement, (D) length of distal resection margin, (E) incompleteness of mesorectum, and (F) harvested lymph nodes. Random-effects models were used for meta-analysis of length of circumferential resection margin and distal resection margin. Fixed-effects models were used for meta-analysis of circumferential resection margin and distal resection margin involvement. Mean differences and risk ratios are shown with 95% confidence intervals (CIs). TaTME, transanal total mesorectal excision; LaTME, laparoscopic total mesorectal excision; SD, standard deviation; IV, inverse variance; df, degree of freedom.

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D		TaTME			LaTME			Mean difference	Mean difference					
Study or subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, random, 95% CI		IV, random, 95	% CI			
Alamili 2019	15	14.5	40	31.25	13.75	20	10.0%	-16.25 [-23.77, -8.73]				-		
Chang 2018	13.5	10.5	23	15.5	10.5	23	10.9%	-2.00 [-8.07, 4.07]						
Chen 2016	24	12	50	15	9	100	12.2%	9.00 [5.24, 12.76]						
Chen 2019	16	14	39	19	13	64	11.3%	-3.00 [-8.43, 2.43]						
Fernández-Hevia 2015	28	18	37	17	13	37	10.2%	11.00 [3.85, 18.15]		-				
Mege 2018	13	9	34	14	12	34	11.5%	-1.00 [-6.04, 4.04]						
Perdawood 2019	33.45	14.5	29	25.41	15.78	29	9.8%	8.04 [0.24, 15.84]				-		
Persiani 2018	25	13.75	46	15	8.75	46	11.7%	10.00 [5.29, 14.71]						
Roodbeen 2019	20	7.5	41	20	7.88	41	12.4%	0.00 [-3.33, 3.33]						
Total (95% CI)			339			394	100.0%	1.88 [-2.96, 6.73]			•			
Heterogeneity: $Tau^2 = 46$	6.42; Chi	² = 62.5	50, df =	= 8 (P <	0.001)	; I ² = 8	7%			+ +		—		
	/								-20 -	-10 0	10	20		

Favours TaTME Favours LaTME

Test for overall effect: Z = 0.76 (P = 0.45)

E	TaTM	٨F	LaTI	ME		Risk ratio		Ri	sk ratio		
Study or subgroup	Events Total		Events	Total	Weight	M-H, fixed, 95% Cl		M-H, f	ixed, 95%	CI	
Alamili 2019	11	40	1	20	3.2%	5.50 [0.76, 39.66]					
Bjoern 2019	6	49	8	36	22.3%	0.55 [0.21, 1.45]			-		
Fernández-Hevia 2015	1	37	0	37	1.2%	3.00 [0.13, 71.34]					
Lelong 2017	0	34	2	38	5.7%	0.22 [0.01, 4.48]			_		
Mege 2018	7	34	4	34	9.7%	1.75 [0.56, 5.43]					
Perdawood 2019	14	29	18	29	43.5%	0.78 [0.49, 1.25]		-	╼┼		
Persiani 2018	2	46	2	46	4.8%	1.00 [0.15, 6.80]			_	_	
Rasulov 2016	4	22	4	23	9.5%	1.05 [0.30, 3.67]			<u> </u>		
Roodbeen 2019	0	41	0	41		Not estimable					
Rubinkiewicz 2019	0	23	0	23		Not estimable					
Veltcamp 2019	0	27	0	27		Not estimable					
Total (95% CI)		382		354	100.0%	1.00 [0.70, 1.45]			•		
Total events	45		39								
Heterogeneity: Chi ² = 7.8	1, df = 7 (P	= 0.35); l	² = 10%				0.01	01	1	10	100
Test for overall effect: Z =	0.03 (P = 0	.98)					Fav	ours TaTME	Favo	ours La	TME

Г	-	ГаТМЕ			LaTME			Mean difference	Mean difference						
Study or subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, fixed, 95% CI		IV,	fixed, 95%	% CI			
Alamili 2019	22	22	40	19.5	6.5	20	3.3%	2.50 [-4.89, 9.89]							
Chang 2018	22.8	10.8	23	19.5	8.6	23	5.6%	3.30 [-2.34, 8.94]							
Chen 2016	16.7	7.8	50	17.4	8.9	100	23.1%	-0.70 [-3.48, 2.08]							
Chen 2019	20.8	9	39	18.8	8.1	64	14.9%	2.00 [-1.45, 5.45]				,			
Lelong 2017	14	7	34	12	5.25	38	21.4%	2.00 [-0.88, 4.88]				·			
Mege 2018	14	10	34	14	8	34	9.6%	0.00 [-4.30, 4.30]			<u> </u>				
Perdawood 2019	26.45	8.7	29	26.69	11.16	29	6.7%	-0.24 [-5.39, 4.91]							
Rasulov 2016	22	13.5	22	26	12	23	3.2%	-4.00 [-11.47, 3.47]	-						
Roodbeen 2019	18	3.25	41	20	12	41	12.3%	-2.00 [-5.81, 1.81]							
Total (95% CI)			312			372	100.0%	0.44 [-0.89, 1.78]							
Heterogeneity: Chi ² = 6.8	88, df = 8	6 (P = 0).55); I ²	= 0%								<u>_</u>			
Test for overall effect: Z	= 0.65 (P	= 0.52	2)						-10	-5	0	5	10		

Fig. 5. Continued.

studies ($I^2 = 0\%$; P = 0.550).

Publication bias

Symmetrical funnel plots of the conversion rate, major



complication (Clavien-Dindo classification III-V), CRM involvement, and incompleteness of mesorectum suggested no publication bias exists in the meta-analysis (Supplementary Fig. 3), and all of the studies were within the 95% CIs.

DISCUSSION

In the present systematic review and meta-analysis, we evaluated the bottom-up approach of TaTME compared to conventional LaTME with respect to perioperative and pathological outcomes. Our meta-analysis shows that TaTME has a significantly lower conversion rate compared to LaTME. Open surgical conversion is one of the major problems in laparoscopic surgery for rectal cancer. Conversion rates in laparoscopic surgery depend on several factors such as patientrelated factors, surgeon-related factors, and procedural factors. Unlike other factors, patient-related factors are beyond the control of the surgeon. Therefore, to determine the success of the laparoscopic approach, it is important to consider patientrelated factors such as sex, obesity, tumor stage, and previous abdominal operation [40-42]. Furthermore, patient selection based on accurate tumor characteristics can reduce conversion rates [43]. TaTME has emerged as a feasible alternative surgical option for conventional laparoscopic surgery, particularly in patients with obesity, narrow pelvis, or deep anterior rectal tumors [44]. Notably, our pooled analysis showed significantly lower conversion rates in patients undergoing TaTME (Fig. 2C), suggesting that the transanal approach can overcome limitations on patient-related factors.

While several studies have reported that the 2-team approach of the TaTME procedure has shorter operation times, but this issue is one of the major controversies currently being raised. In the studies included in our meta-analysis, a 2-team approach [24,35], a 1-team approach [28,30,32,34,36,38], and both approaches [25,27,31] were performed. The 2-team approach tended to have shorter operation times compared to the 1-team approach, but there were not enough literatures available for subgroup analysis. TaTME's approach may be carried out differently depending on the circumstances, such as the national healthcare system and health insurance policies or the status of medical personnel in hospitals. In this regard, a recent study reported cost analysis, including surgical supplies along with surgical outcomes [45]. When discussing TaTME's approach in the future, surgical outcomes such as operation time and cost analysis also need to be considered.

Anastomotic leakage is a major complication of colorectal surgery causing increased mortality, the incidence of which has persisted over the last years [46,47]. Recently, an international multicenter cohort study reported that TaTME had a higher risk of anastomotic leakage than LaTME [13]. On the other hand, a recent cohort study of the incidence of anastomosis and learning curves after TaTME reported that surgeon proficiency can have a positive effect on the occurrence of anastomotic leakage. Our findings on the perioperative outcomes showed that LaTME tended to have a higher rate of anastomotic leakage than TaTME, but there was no statistically significant difference (Fig. 4B). Since a wide range of anastomotic leak rates depends on the definition, clinical setting, and follow-up period [29], subgroup analysis of the related factors needs to be conducted through large-scale study to draw clear conclusions.

A Norwegian research group noted that TaTME could increase the risk of local recurrence [48]. The results showed local recurrence of a new pattern characterized by rapid, multifocal growth in the pelvic cavity and sidewalls early after TaTME, which led to a nationwide discontinuance of TaTME [48,49]. It should be noted that although there are concerns about the risk of local recurrence, conclusions cannot be drawn due to insufficient numbers and lack of long-term oncological followup from the national multicenter study. These results further suggest that structured training with proctorship experienced proponents is essential.

One of the potential benefits of TaTME is improved specimen quality, defined as CRM involvement, DRM distance, and mesorectal completeness. These pathological outcomes are the most potent prognostic factors predicting local recurrence [50,51]. Fifteen studies involving 1,863 patients showed a tendency to a reduced CRM positive status of TaTME in comparison with LaTME (Fig. 5A). Regarding the DRM distance, 4 of 9 studies reported that TaTME showed a longer DRM compared to LaTME [24,25,31,35] (Fig. 5D). Furthermore, incompleteness of mesorectum in TaTME was similar to the results from LaTME (Fig. 5E). However, contrary to the results reported in the individual studies, meta-analysis showed TaTME did not lead to significant improvement in pathological outcomes when compared to LaTME.

Among the surgical outcomes of TaTME and LaTME, studies evaluating the length of CRM and DRM showed considerable heterogeneity, and there is one possible explanation for the significant heterogeneity. We used the method of estimating the median value as the mean value for meta-analysis, and it seems that heterogeneity occurred in this process. In the study of Alamili et al. [32], the median was similar between the 2 groups, but the difference in sample size resulted in a large difference in the estimate of the mean value. The main limitation of this method is that the outcome distribution is assumed to be normal, and new methods have been developed to overcome this limitation [52]. However, since any method may make an unsuitable assumption, results should be interpreted with additional explanations.

In conclusion, this updated systematic review and metaanalysis showed that TaTME has potential advantages in surgical outcomes, suggesting that TaTME can be performed as a New Health Technology for rectal cancer patients in South Korea. However, since this study was conducted based on limited evidence, the results of a well-designed multicenter RCTs need to be further evaluated to apply our conclusions extensively; and expertise and training to safely perform



TaTME should also be considered.

SUPPLEMENTARY MATERIALS

Supplementary Table 1 and Supplementary Fig. 1–3 can be found *via* https://doi.org/10.4174/astr.2021.101.3.167

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Conflict of Interest

NKK served as a chairman of TaTME Assessment Committee. SHK, DHL, JHB, SSC, JYS, and CSE as members of the Committee, reviewed the surgical technology in depth for nHTA. No other potential conflict of interest relevant to this article was reported.

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Conceptualization, Investigation: SHK, YIJ Formal Analysis: SHK Methodology: SHK, DHL, JHB, SSC, JYS, CSE Project Administration: NKK Writing – Original Draft: SHK Writing – Review & Editing: All authors

REFERENCES -

- Bray F, Ferlay J, Soerjomataram I, Siegel RL, Torre LA, Jemal A. Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. CA Cancer J Clin 2018;68:394-424.
- Heald RJ, Husband EM, Ryall RD. The mesorectum in rectal cancer surgerythe clue to pelvic recurrence?. Br J Surg 1982;69:613-6.
- 3. Heald RJ, Moran BJ, Ryall RD, Sexton R, MacFarlane JK. Rectal cancer: the Basingstoke experience of total mesorectal excision, 1978-1997. Arch Surg 1998;133:894-9.
- 4. Veldkamp R, Kuhry E, Hop WC, Jeekel J, Kazemier G, Bonjer HJ, et al. Laparoscopic surgery versus open surgery for colon cancer: short-term outcomes of a randomised trial. Lancet Oncol 2005;6:477-84.
- 5. Guillou PJ, Quirke P, Thorpe H, Walker J, Jayne DG, Smith AM, et al. Shortterm endpoints of conventional versus laparoscopic-assisted surgery in patients

with colorectal cancer (MRC CLASICC trial): multicentre, randomised controlled trial. Lancet 2005:365:1718-26.

- Penninckx F, Kartheuser A, Van de Stadt J, Pattyn P, Mansvelt B, Bertrand C, et al. Outcome following laparoscopic and open total mesorectal excision for rectal cancer. Br J Surg 2013;100:1368-75.
- Lujan J, Valero G, Hernandez Q, Sanchez A, Frutos MD, Parrilla P. Randomized clinical trial comparing laparoscopic and open surgery in patients with rectal cancer. Br J Surg 2009;96:982-9.
- Mizrahi I, Sands DR. Transanal total mesorectal excision for rectal cancer: a review. Ann Laparosc Endosc Surg 2017;2:144.
- Sylla P, Rattner DW, Delgado S, Lacy AM. NOTES transanal rectal cancer resection using transanal endoscopic microsurgery and laparoscopic assistance. Surg Endosc 2010;24:1205-10.
- Penna M, Hompes R, Arnold S, Wynn G, Austin R, Warusavitarne J, et al. Transanal total mesorectal excision: international

registry results of the first 720 cases. Ann Surg 2017;266:111-7.

- 11. de Lacy FB, van Laarhoven JJ, Pena R, Arroyave MC, Bravo R, Cuatrecasas M, et al. Transanal total mesorectal excision: pathological results of 186 patients with mid and low rectal cancer. Surg Endosc 2018:32:2442-7.
- 12. Abbott SC, Stevenson AR, Bell SW, Clark D, Merrie A, Hayes J, et al. An assessment of an Australasian pathway for the introduction of transanal total mesorectal excision (taTME). Colorectal Dis 2018;20:O1-6.
- 13. 2017 European Society of Coloproctology (ESCP) collaborating group. An international multicentre prospective audit of elective rectal cancer surgery; operative approach versus outcome, including transanal total mesorectal excision (TaTME). Colorectal Dis 2018;20 Suppl 6:33-46.
- Wasmuth HH, Faerden AE, Myklebust TÅ, Pfeffer F, Norderval S, Riis R, et al. Transanal total mesorectal excision for

rectal cancer has been suspended in Norway. Br J Surg 2020;107:121-30.

- 15. Sylla P, Knol JJ, D'Andrea AP, Perez RO, Atallah SB, Penna M, et al. Urethral injury and other urologic injuries during transanal total mesorectal excision: an international collaborative study. Ann Surg 2021:274:e115-25.
- Caycedo-Marulanda A, Verschoor CP. Experience beyond the learning curve of transanal total mesorectal excision (taTME) and its effect on the incidence of anastomotic leak. Tech Coloproctol 2020;24:309-16.
- Persiani R, Agnes A, Belia F, D'Ugo D, Biondi A. The learning curve of TaTME for mid-low rectal cancer: a comprehensive analysis from a five-year institutional experience. Surg Endosc 2020 Oct 26 [Epub]. https://doi.org/10.1007/s00464-020-08115-0.
- 18. Liberati A, Altman DG, Tetzlaff J, Mulrow C, Gøtzsche PC, Ioannidis JP, et al. The PRISMA statement for reporting systematic reviews and meta-analyses of studies that evaluate healthcare interventions: explanation and elaboration. BMJ 2009:339:b2700.
- 19. Stang A. Critical evaluation of the Newcastle-Ottawa scale for the assessment of the quality of nonrandomized studies in meta-analyses. Eur J Epidemiol 2010;25:603-5.
- 20. Quirke P, Steele R, Monson J, Grieve R, Khanna S, Couture J, et al. Effect of the plane of surgery achieved on local recurrence in patients with operable rectal cancer: a prospective study using data from the MRC CR07 and NCIC-CTG CO16 randomised clinical trial. Lancet 2009;373:821-8.
- 21. Efthimiou O. Practical guide to the metaanalysis of rare events. Evid Based Ment Health 2018;21:72-6.
- 22. Hozo SP, Djulbegovic B, Hozo I. Estimating the mean and variance from the median, range, and the size of a sample. BMC Med Res Methodol 2005;5:13.
- 23. Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. BMJ 2003:327:557-60.

- 24. Fernández-Hevia M, Delgado S, Castells A, Tasende M, Momblan D, Díaz del Gobbo G, et al. Transanal total mesorectal excision in rectal cancer: short-term outcomes in comparison with laparoscopic surgery. Ann Surg 2015:261:221-7.
- 25. Chen CC, Lai YL, Jiang JK, Chu CH, Huang IP, Chen WS, et al. Transanal total mesorectal excision versus laparoscopic surgery for rectal cancer receiving neoadjuvant chemoradiation: a matched case-control study. Ann Surg Oncol 2016;23:1169-76.
- 26. Rasulov AO, Mamedli ZZ, Gordeyev SS, Kozlov NA, Dzhumabaev HE. Short-term outcomes after transanal and laparoscopic total mesorectal excision for rectal cancer. Tech Coloproctol 2016;20:227-34.
- 27. Lelong B, Meillat H, Zemmour C, Poizat F, Ewald J, Mege D, et al. Short- and midterm outcomes after endoscopic transanal or laparoscopic transabdominal total mesorectal excision for low rectal cancer: a single institutional case-control study. J Am Coll Surg 2017;224:917-25.
- 28. Chang TC, Kiu KT. Transanal total mesorectal excision in lower rectal cancer: comparison of short-term outcomes with conventional laparoscopic total mesorectal excision. J Laparoendosc Adv Surg Tech A 2018;28:365-9.
- 29. Detering R, Roodbeen SX, van Oostendorp SE, Dekker JT, Sietses C, Bemelman WA, et al. Three-year nationwide experience with transanal total mesorectal excision for rectal cancer in the Netherlands: a propensity score-matched comparison with conventional laparoscopic total mesorectal excision. J Am Coll Surg 2019;228:235-44.
- 30. Mege D, Hain E, Lakkis Z, Maggiori L, Prost À la Denise J, Panis Y. Is trans-anal total mesorectal excision really safe and better than laparoscopic total mesorectal excision with a perineal approach first in patients with low rectal cancer?: a learning curve with case-matched study in 68 patients. Colorectal Dis 2018;20:O143-51.
- Persiani R, Biondi A, Pennestrì F, Fico V, De Simone V, Tirelli F, et al. Transanal

total mesorectal excision vs laparoscopic total mesorectal excision in the treatment of low and middle rectal cancer: a propensity score matching analysis. Dis Colon Rectum 2018;61:809-16.

- 32. Alamili M, Levic K, Kanstrup K, Bisgaard T, Bulut O. Inflammatory response after transanal total mesorectal excision. Dan Med J 2019:66:A5555.
- 33. Bjoern MX, Nielsen S, Perdawood SK. Quality of life after surgery for rectal cancer: a comparison of functional outcomes after transanal and laparoscopic approaches. J Gastrointest Surg 2019;23:1623-30.
- 34. Chen YT, Kiu KT, Yen MH, Chang TC. Comparison of the short-term outcomes in lower rectal cancer using three different surgical techniques: transanal total mesorectal excision (TME), laparoscopic TME, and open TME. Asian J Surg 2019:42:674-80.
- 35. Perdawood SK, Warnecke M, Bjoern MX, Eiholm S. The pattern of defects in mesorectal specimens: is there a difference between transanal and laparoscopic approaches? Scand J Surg 2019;108:49-54.
- 36. Roodbeen SX, Penna M, Mackenzie H, Kusters M, Slater A, Jones OM, et al. Transanal total mesorectal excision (TaTME) versus laparoscopic TME for MRI-defined low rectal cancer: a propensity score-matched analysis of oncological outcomes. Surg Endosc 2019;33:2459-67.
- 37. Rubinkiewicz M, Zarzycki P, Witowski J, Pisarska M, Gajewska N, Torbicz G, et al. Functional outcomes after resections for low rectal tumors: comparison of transanal with laparoscopic total mesorectal excision. BMC Surg 2019;19:79.
- 38. Sparreboom CL, Komen N, Rizopoulos D, van Westreenen HL, Doornebosch PG, Dekker JW, et al. Transanal total mesorectal excision: how are we doing so far? Colorectal Dis 2019;21:767-74.
- 39. Veltcamp Helbach M, Koedam TW, Knol JJ, Velthuis S, Bonjer HJ, Tuynman JB, et al. Quality of life after rectal cancer surgery: differences between laparoscopic

and transanal total mesorectal excision. Surg Endosc 2019;33:79-87.

- 40. Thorpe H, Jayne DG, Guillou PJ, Quirke P, Copeland J, Brown JM, et al. Patient factors influencing conversion from laparoscopically assisted to open surgery for colorectal cancer. Br J Surg 2008;95:199-205.
- 41. Pandya S. Murray JJ. Coller JA. Rusin LC. Laparoscopic colectomy: indications for conversion to laparotomy. Arch Surg 1999:134:471-5.
- 42. Pikarsky AJ, Saida Y, Yamaguchi T, Martinez S, Chen W, Weiss EG, et al. Is obesity a high-risk factor for laparoscopic colorectal surgery? Surg Endosc 2002;16:855-8.
- 43. Veldkamp R, Gholghesaei M, Bonjer HJ, Meijer DW, Buunen M, Jeekel J, et al. Laparoscopic resection of colon cancer: consensus of the European Association of Endoscopic Surgery (EAES). Surg Endosc 2004;18:1163-85.
- 44. Rouanet P, Mourregot A, Azar CC, Carrere

S, Gutowski M, Quenet F, et al. Transanal endoscopic proctectomy: an innovative procedure for difficult resection of rectal tumors in men with narrow pelvis. Dis Colon Rectum 2013;56:408-15.

- 45. Di Candido F, Carvello M, Keller DS, Vanni E, Maroli A, Montroni I, et al. A comparative cost analysis of transanal and laparoscopic total mesorectal excision for rectal cancer. Updates Surg 2021:73:85-91.
- 46. Krarup PM, Jorgensen LN, Andreasen AH, Harling H: Danish Colorectal Cancer Group. A nationwide study on anastomotic leakage after colonic cancer surgery. Colorectal Dis 2012;14:e661-7.
- Paun BC, Cassie S, MacLean AR, Dixon E, Buie WD. Postoperative complications following surgery for rectal cancer. Ann Surg 2010;251:807-18.
- Larsen SG, Pfeffer F, Kørner H; Norwegian Colorectal Cancer Group. Norwegian moratorium on transanal total mesorectal excision. Br J Surg 2019;106:1120-1.

- 49. Gachabayov M, Bergamaschi R. Is taTME delivering? Updates Surg 2019;71:13-5.
- 50. Nagtegaal ID, Marijnen CA, Kranenbarg EK, van de Velde CJ, van Krieken JH; Pathology Review Committee, et al. Circumferential margin involvement is still an important predictor of local recurrence in rectal carcinoma: not one millimeter but two millimeters is the limit. Am J Surg Pathol 2002;26:350-7.
- 51. Lin HH, Lin JK, Lin CC, Lan YT, Wang HS, Yang SH, et al. Circumferential margin plays an independent impact on the outcome of rectal cancer patients receiving curative total mesorectal excision. Am J Surg 2013;206:771-7.
- 52. McGrath S, Zhao X, Steele R, Thombs BD, Benedetti A: DEPRESsion Screening Data (DEPRESSD) Collaboration. Estimating the sample mean and standard deviation from commonly reported quantiles in meta-analysis. Stat Methods Med Res 2020 Jan 30 [Epub]. https://doi. org/10.1177/0962280219889080.