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Comparison of the Effects of Ezetimibe-Statin Combination Therapy on Major Adverse Cardiovascular Events in Patients with and without Diabetes: A Meta-Analysis

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Background: Ezetimibe-statin combination therapy has been found to reduce low density lipoprotein cholesterol levels and the risk of major adverse cardiovascular events (MACEs) in large trials. We sought to examine the differential effect of ezetimibe on MAC-Es when added to statins according to the presence of diabetes.

Methods: Randomized clinical trials with a sample size of at least 50 participants and at least 24 weeks of follow-up that compared ezetimibe-statin combination therapy with a statin- or placebo-controlled arm and reported at least one MACE, stratified by diabetes status, were included in the meta-analysis and meta-regression.

Results: A total of seven trials with 28,191 enrolled patients (mean age, 63.6 years; 75.1% men; 7,298 with diabetes [25.9%]; mean follow-up, 5 years) were analysed. MACEs stratified by diabetes were obtained from the published data (two trials) or through direct contact (five trials). No significant heterogeneity was observed among studies (I^2 =14.7%, P=0.293). Ezetimibe was associated with a greater reduction of MACE risk in subjects with diabetes than in those without diabetes (pooled relative risk, 0.84 vs. 0.93; $P_{heterogeneity}$ =0.012). In the meta-regression analysis, the presence of diabetes was associated with a greater reduction of MACE risk when ezetimibe was added to statins (β =0.87, P=0.038).

Conclusion: Ezetimibe-statin combination therapy was associated with greater cardiovascular benefits in patients with diabetes than in those without diabetes. Our findings suggest that ezetimibe-statin combination therapy might be a useful strategy in patients with diabetes at a residual risk of MACEs.

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Keywords: Ezetimibe; Myocardial infarction; Stroke; Hydroxymethylglutaryl-CoA reductase inhibitors; Diabetes mellitus

INTRODUCTION

Statins (3-hydroxy-3-methylglutaryl coenzyme A reductase inhibitors) have shown efficacy in lowering cholesterol levels and reducing the risk of cardiovascular events in the setting of primary and secondary prevention [1,2]. Given the major contribution of cardiovascular events to morbidity and mortality in patients with diabetes, high-intensity statins are recommended for patients with diabetes [3,4]. However, individuals with diabetes have substantial residual cardiovascular risk, even when receiving statin therapy, leading to an unmet need for additional lipidmodifying strategies [5].

Ezetimibe, a Niemann-Pick C1-like1 (NPC1L1) inhibitor, blocks intestinal cholesterol absorption, leading to the reduction of circulating cholesterol levels via a distinct mechanism from that of statins [6,7]. A large randomized controlled trial (RCT), the Improved Reduction of Outcomes: Vytorin Efficacy International Trial (IMPROVE-IT), demonstrated the efficacy of ezetimibe-statin combination therapy on the reduction of cholesterol levels and major cardiovascular adverse events (MACEs) in patients who had recently experienced a myocardial infarction [8]. Notably, in a subgroup analysis, the beneficial effect of ezetimibe added to statins on MACEs was more prominent in patients with diabetes than in patients without diabetes [8]. Results from another large, placebo-controlled trial investigating the efficacy of ezetimibe-statin combination therapy in reducing cardiovascular events in chronic kidney disease patients also found a similar preferential effect of ezetimibe-statin combination therapy in patients with diabetes [9]. Given the high residual cardiovascular risk in patients with diabetes who are receiving treatment, these findings suggest that ezetimibe might provide additional benefits for preventing cardiovascular events, particularly in patients with diabetes. However, this potential differential effect of ezetimibe according to presence of diabetes has not been assessed as a primary outcome in pooled results from RCTs.

In this meta-analysis, we compared the effect of ezetimibestatin combination therapy on MACEs to that of statins alone or placebo in patients with and without diabetes, based on the pooled results of RCTs.

METHODS

Data sources, search strategy, and selection criteria

We conducted a meta-analysis based on the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines [10]. Relevant studies were identified by searching the following data sources: MEDLINE via PubMed, Embase, and the Central Controlled Trials Register of the Cochrane Collaboration (from 1994 to December 2016). The following text words and medical subject headings were used without language restriction: "ezetimibe," "ezetimibe-simvastatin drug combination," "simvastatin," "pravastatin," "lovastatin," "atorvastatin," "rosuvastatin," "fluvastatin," "pitavastatin," and "hydroxymethylglutaryl-CoA reductase inhibitors." The reference lists of identified studies were also scanned to find potentially relevant studies. Two independent authors (Y.H.L and N.H.) performed the literature search, data extraction, and quality assessment with a standardized method, and a third reviewer (E.S.K.) adjudicated any discrepancies. Quality assessment was done using the Cochrane Collaboration's tool for assessing the risk of bias in randomized trials (Supplemental Fig. S1) [11]. This study was approved by Institutional Review Board of Severance Hospital, Yonsei University (no. 4-2015-0637).

Data extraction and quality assessment

Standard information was extracted from published reports and unpublished data obtained from investigators into a spreadsheet. We requested and received data using a formal question sheet for trials with unpublished information. We collected data on the number of randomized patients and the occurrence of MAC-Es in each ezetimibe and comparator group in the overall participants, as well as in subgroups divided by the presence of diabetes. Mean age, body mass index, follow-up duration, and the difference in the decrease of serum low density lipoprotein cholesterol (LDL-C) concentration between the ezetimibe and control groups during the study were tabulated for each study.

Statistical analysis

Relative risks (RRs) and 95% confidence intervals (CIs) were calculated from the event numbers, with the total number of patients as the denominator for individual studies. Heterogeneity across studies was estimated using the I^2 statistic [12]. I^2 values

ranging from 0% to 40% were regarded as indicating no important heterogeneity; moderate, substantial, and considerable heterogeneity were defined as I^2 values ranging from 30% to 60%, 50% to 90%, and 75% to 100%, respectively [13]. Weighted pooled treatment effects were obtained with a random-effects model to provide a more conservative assessment of the average effect size. The heterogeneity of the pooled effect between subgroups was calculated using the Cochran Q statistic, with the following formula: $Q = \sum [(1/variance of individual study) \times (ef$ fect of individual study-effect of pooled study)]², where variance of individual study=[(upper limit–lower limit)/ $(2 \times z)$]² [14]. A funnel plot with symmetry testing by the Egger linear regression method was used to test for potential publication bias [15]. Sensitivity analyses were performed by repeating analyses while removing one study at a time using the 'metaninf' command (STATA). Analyses confined to statin-controlled trials were also performed, with the exclusion of placebo-controlled trials. Random-effects meta-regression models with inverse variance weighting were built to assess whether the presence of diabetes explained the variance in the estimated RR for MACEs observed between trials. Two-sided P values < 0.05 were considered to indicate statistical significance. All statistical analyses were performed with STATA version 14.0 (Stata Corp., College Station, TX, USA).

RESULTS

Characteristics of trials

Among the 13,220 identified records, 347 randomized, placebo or statin-controlled endpoint trials of ezetimibe were screened (Fig. 1). Studies were included if they were completed RCTs comparing the effects of adding ezetimibe to any statin or placebo on the incidence of MACEs and if they reported the clinical outcomes in participants stratified by the presence of diabetes. We also contacted investigators from eight potentially relevant trials about unpublished data for incident MACEs in participants stratified by diabetes, and received and included data from five of those trials. Finally, a total of seven studies, two with published data [8,9] and five with previously unpublished data that had not been analysed until our request [16-20], were included in the meta-analysis. The included studies enrolled 28,191 patients (7,298 with diabetes [25.9%]) with stable angina, recent acute coronary syndrome, chronic kidney disease, peripheral arterial occlusive disease, or hypercholesterolemia (Table 1). The mean age of study subjects was 63.6 years and 75.1% were men. The mean follow-up duration of the studies



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Fig. 1. Flow diagram of the literature search to identify randomized controlled trials (RCTs) comparing the differential effect of ezetimibe combination therapy on the reduction of major adverse cardiovascular events (MACEs) according to the presence of diabetes. DM, diabetes mellitus.

was approximately 5 years, according to the weighted average. The prevalence of diabetes varied from 22.6% to 49.7%. Only one study, the Study of Heart and Renal Protection (SHARP) trial, was a placebo-controlled study (vs. an ezetimibe-simvastatin combination), whereas other trials included statin users as the control group. A greater LDL-C reduction (%) was shown in the ezetimibe and statin combination group than in the control group (statins or placebo), regardless of differences in the intensity and doses in the statin-controlled trials.

Outcome analysis

Fig. 2 shows the pooled association of ezetimibe combination therapy with MACE risk according to the presence of diabetes. No significant heterogeneity was observed across the trials ($I^2=14.7\%$, P=0.293). In the included patients, a total of 6,581 MACEs occurred during follow-up. The definitions of MACEs were generally consistent among studies. Fig. 2A shows that the association of ezetimibe combination therapy with a lower MACE risk was greater in the pooled RR from subgroups with diabetes (RR, 0.84; 95% CI, 0.77 to 0.91) than in the pooled RR

Table 1. I MACEs G	Data of Subject rouped by the	cts in Seven Presence of	Randor Diabete	nized C SS	ontrolle	ed Trials of	f Ezetimib	e Combinat	tion Therap	y with St	atin or Placebo Comparatc	or Arms that Rej	ported Incident
Study	DM/Total, no. (%)	Target population	Mean age, yr	Men, %	Mean BMI, kg/m²	LDL-C reduction in treatment group, %	LDL-C reduction in control group, %	Treatment ^a	Control	Median follow-up wk	. MACE definition	MACEs in treatment group, no. (%)	MACEs in control group, no. (%)
West	29/67 (43.2)	PAOD	63.5	55.9	29.0	-42.8	-26.3	S40+E10	S40 or previous statin	96	CV death, non-fatal MI, ischemic stroke, and TIA	16/51 (31)	6/16 (38)
SHARP	2,094/9,270 (22.6)	CKD	62.0	62.6	27.1	-35.6	-2.4	S20+E10	Placebo	240	CV death, non-fatal MI, ischemic stroke, coronary revascularization	526/4,650 (11)	619/4,620(13)
Kouvelos	79/262 (30.2)	Elective vascular surgery	71.0	89.7	NA	-48.8	-39.0	R10+E10	R10	48	CV death, non-fatal MI, ischemic stroke, hospital- ization for USA	9/126 (7)	17/136 (13)
IM- Prove- IT	4,933/18,144 (27.2)	ACS	63.6	75.7	28.3	-42.9	-26.1	S40+E10	S40	288	CV death, non-fatal MI, ischemic stroke, hospital- ization for USA, coronary revascularization	2,572/9,067 (28)	2,742/9,077 (30)
Suzuki	78/157 (49.7)	Hypercho- lesterol- emia	64.0	64.0	25.5	-15.0	-14.3	Any statin+E10	Any statin	144	CV death, non-fatal MI, ischemic stroke, hospital- ization for USA, coronary revascularization	1/86(1)	4/71 (6)
PRECISE- IVUS	60/202 (29.7)	ACS, stable angina	66.5	0.07	25.5	-19.3	4.3	A10+E10	A10	48	CV death, non-fatal MI, ischemic stroke, hospital- ization for USA, coronary revascularization	18/100 (18)	25/102 (25)
HEAVEN	25/89 (25.0)	Stable angina	64.3	71.9	NA	-35.5	-3.7	A80+E10	Any statin	48	CV death, non-fatal MI, ischemic stroke, hospital- ization for USA, coronary revascularization	13/42 (31)	13/47 (28)
MACE, ma vascular; M IMPROVE Inhibitor or tion study. *A10 atorva	jor adverse carc II, myocardial ii -IT, the Improvi- Synthesis Inhit statin (10 mg), statin (10 mg),	liovascular evu nfarction; TIA ed Reduction vitor Evaluated S20 simvastati	ent; DM, , transier of Outco 1 by Intra in (20 mg	, diabetes it ischer innes: Vy ivascular g), S40 si	s mellitus nic attack torin Eff Ultrasou imvastat	s; BMI, body c; SHARP, tl icacy Intern und Study; F in (40 mg), I	/ mass inde he Study of ational Tria HEAVEN, V A10 rosuva	x; LDL-C, lo 'Heart and Re J; ACS, acute /irtual histolo statin (10 mg)	w density lip enal Protectic e coronary sy gy evaluation), E10 ezetim	oprotein cl on; CKD, c ndrome; P 1 of athero ibe (10 mg	nolesterol; PAOD, peripheral a chronic kidney disease; NA, nc RECISE-IVUS, Plaque Regre sclerosis regression during ato ().	tfery occlusive dis ot available; USA, ssion With Choles rvastatin and ezeti	ease; CV, cardio- unstable angina; tterol Absorption mibe administra-

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from subgroups without diabetes (RR, 0.93; 95% CI, 0.85 to 1.02; $P_{heterogeneity}$ =0.012). A similar result was observed when the placebo-controlled study (SHARP) was excluded from the pooled analysis (RR, 0.86; 95% CI, 0.78 to 0.94 vs. RR, 0.97; 95% CI, 0.92 to 1.03; $P_{heterogeneity}$ =0.022) (Fig. 2B), indicating a

Study	RR (9	5% CI)	% weight	Sample size
Subgroups without diabetes				
West-NonDM	1.09(0.	27, 4.33)	0.24	47
SHARP-NonDM	 ➡ 0.86 (0. 	74, 0.99)	16.12	7,176
HEAVEN-NonDM	1.04 (0.	55, 1.97)	1.13	64
Suzuki-NonDM <	0.29 (0.	01, 7.00)	0.04	79
Kouvelos-NonDM	0.43 (0.	16, 1.17)	0.47	183
PRECISE-IVUS-NonDM	0.69 (0.	38, 1.28)	1.25	142
IMPROVE-IT-NonDM	• 0.98 (0.	92, 1.04)	40.02	13,211
Subtotal (1-squared = 12.5% , $P=0.334$) 0.93 (0.	85, 1.02)	59.28	
Subgroups with diabetes				
West-DM	0.72 (0.	32, 1.62)	0.71	20
SHARP-DM	- 0.78 (0.	(0.94)	10.46	2,094
HEAVEN-DIVI	0.90(0.	57, 2.20)	0.59	25
Suzuki-Divi ———	0.26(0.	(3, 2.37)	0.10	70
PRECISE-IVUS-DM	0.78(0.	25, 2.09)	0.31	60
IMPROVE-IT-DM	0.85 (0.	78 0 94)	28.16	4 933
Subtotal (I-squared= $0.0\% P=0.910$)	0.86 (0.	77 0 91)	40.72	1,755
50510101 (1 September 0.070,1 0.910)	v 0.01(0.	,,,0.51)	10.72	
Overall (I-squared=14.7%, P=0.293)	0.89 (0.	83, 0.95)	100.00	
	0.2 1 5			A
Study	PP (0	5% CD 1	%	Sample
Study	RR (9	5% CI)	% weight	Sample size
Study Subgroups without diabetes	RR (9.	5% CI)	% weight	Sample
Study Subgroups without diabetes West-NonDM	RR (9	5% CI)	% weight 0.15	Sample size
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM	RR (9	5% CI) 27, 4.33) 55, 1.97)	% weight 0.15 0.70	Sample size 47 64
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM	RR (9.	5% CI) 27, 4.33) 55, 1.97) 01, 7.00)	% weight 0.15 0.70 0.03	Sample size 47 64 79
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Kouvelos-NonDM	RR (9.	5% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17)	% weight 0.15 0.70 0.03 0.29	Sample size 47 64 79 183
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Kouvelos-NonDM PRECISE-IVUS-NonDM	RR (9.	5% CI) v 27,4.33) 55,1.97) 01,7.00) 16,1.17) 38,1.28)	% weight 0.15 0.70 0.03 0.29 0.78	Sample size 47 64 79 183 142
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Kouvelos-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM	RR (9 1.09 (0. 1.04 (0. 0.29 (0. 0.43 (0. 0.69 (0. 0.69 (0. 0.98 (0.	5% CI) x 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04)	% weight 0.15 0.70 0.03 0.29 0.78 65.93	Sample size 47 64 79 183 142 13,211
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Kouvelos-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485)	RR (9 1.09 (0. 1.04 (0. 0.29 (0. 0.43 (0. 0.69 (0. 0.98 (0. 0.97 (0.	5% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88	Sample size 47 64 79 183 142 13,211
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Kouvelos-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes	RR (9	5% CI) x 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88	Sample size 47 64 79 183 142 13,211
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Kouvelos-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM	RR (9 1.09 (0. 1.04 (0. 0.29 (0. 0.43 (0. 0.69 (0. 0.98 (0. 0.97 (0. 0.72 (0.	5% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03) 32, 1.62)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44	Sample size 47 64 79 183 142 13,211 20
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM VEXUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM	RR (9.	5% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03) 32, 1.62) 37, 2.20)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36	Sample size 47 64 79 183 142 13,211 20 25
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Vest-NonDM MPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM Suzuki-DM	RR (9.	55% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03) 32, 1.62) 37, 2.20) 03, 2.37)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36 0.06	Sample size 47 64 79 183 142 13,211 20 25 78
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Kouvelos-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM Suzuki-DM Kouvelos-DM	RR (9 1.09 (0. 1.04 (0. 0.29 (0. 0.43 (0. 0.43 (0. 0.98 (0. 0.97 (0. 0.99 (0. 0.90 (0.	55% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 02, 1.03) 32, 1.62) 37, 2.20) 03, 2.37) 23, 2.60)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36 0.06 0.19	Sample size 47 64 79 183 142 13,211 20 25 78 79
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM VECUSE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM Suzuki-DM Kouvelos-DM PRECISE-IVUIS-DM	RR (9 1.09 (0. 1.04 (0. 0.29 (0. 0.43 (0. 0.69 (0. 0.98 (0. 0.97 (0. 0.97 (0. 0.90 (0. 0.72 (0. 0.90 (0. 0.26 (0. 0.78 (0. 0. 0.78 (0. 0.78 (0. 0. 0. 0. 0. 0. 0. 0. 0. 0.	5% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03) 32, 1.62) 37, 2.20) 03, 2.37) 23, 2.69) 30, 2.61)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36 0.06 0.19 0.24	Sample size 47 64 79 183 142 13,211 20 25 78 79 60
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM VEXUS-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM Suzuki-DM Kouvelos-DM PRECISE-IVUS-DM MRPOVE IT DM	RR (9.	5% CI) 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03) 32, 1.62) 37, 2.20) 03, 2.37) 23, 2.69) 30, 2.61) 78, 0.94)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36 0.06 0.19 0.19 0.29	Sample size 47 64 79 183 142 13,211 20 25 78 79 60 0 4 033
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Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM VEXUS-NonDM PRECISE-IVUS-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM Suzuki-DM Kouvelos-DM PRECISE-IVUS-DM IMPROVE-IT-DM Subtotal (I-squared=0.0%, P=0.928)	RR (9.	5% CI) x 27, 4.33) 55, 1.97) 01, 7.00) 16, 1.17) 38, 1.28) 92, 1.04) 92, 1.03) 32, 1.62) 37, 2.20) 03, 2.37) 23, 2.69) 30, 2.61) 78, 0.94) 78, 0.94)	% 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36 0.06 0.19 0.24 30.83 32.12	Sample size 47 64 79 183 142 13,211 20 25 78 79 60 4,933
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Kouvelos-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM Suzuki-DM Kouvelos-DM PRECISE-IVUS-DM IMPROVE-IT-DM Subtotal (I-squared=0.0%, P=0.928) Overall (I-squared=0.9%, P=0.435)	RR (9. 1.09 (0. 1.04 (0. 0.29 (0. 0.43 (0. 0.69 (0. 0.98 (0. 0.90 (0. 0.98 (0. 0.99 (0. 0.99 (0. 0.98 (0. 0.98 (0. 0.99 (0. 0.99 (0. 0.98 (0. 0.98 (0. 0.99 (0. 0.99 (0. 0.98 (0. 0.99 (0. 0.99 (0. 0.98 (0. 0.99 (0.	5% CI) x 27,4.33) 55,1.97) 01,7.00) 16,1.17) 38,1.28) 92,1.04) 92,1.03) 32,1.62) 37,2.20) 03,2.37) 23,2.69) 30,2.61) 78,0.94) 78,0.94) 89,0.99)	% 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36 0.06 0.19 0.24 30.83 32.12 100.00	Sample size 47 64 79 183 142 13,211 20 25 78 79 60 4,933
Study Subgroups without diabetes West-NonDM HEAVEN-NonDM Suzuki-NonDM Vest-NonDM PRECISE-IVUS-NonDM IMPROVE-IT-NonDM Subtotal (I-squared=0.0%, P=0.485) Subgroups with diabetes West-DM HEAVEN-DM Suzuki-DM Kouvelos-DM PRECISE-IVUS-DM IMPROVE-IT-DM Subtotal (I-squared=0.0%, P=0.928) Overall (I-squared=0.9%, P=0.435)	RR (9 1.09 (0. 1.04 (0. 0.29 (0. 0.43 (0. 0.69 (0. 0.98 (0. 0.98 (0. 0.98 (0. 0.98 (0. 0.86 (0. 0.86 (0. 0.86 (0. 0.86 (0. 0.93 (0. 0. 0.93 (0. 0.93 (0. 0. 0. 0. 0. 0. 0. 0. 0. 0.	5% CI) x 27,4.33) 55,1.97) 01,7.00) 16,1.17) 38,1.28) 92,1.04) 92,1.03) 32,1.62) 37,2.20) 03,2.37) 23,2.69) 30,2.61) 78,0.94) 78,0.94) 89,0.99)	% weight 0.15 0.70 0.03 0.29 0.78 65.93 67.88 0.44 0.36 0.06 0.19 0.24 30.83 32.12 100.00	Sample size 47 64 79 183 142 13,211 20 25 78 79 60 4,933

Fig. 2. Pooled effects of ezetimibe-statin combination therapy on major adverse cardiovascular events grouped by the presence of diabetes within studies. The test for heterogeneity between subgroups was significant (A) in all studies (P=0.012) and (B) after excluding the placebo-controlled trial (SHARP) (P=0.022). RR, risk ratio; CI, confidence interval; DM, diabetes mellitus; SHARP, the Study of Heart and Renal Protection; HEAVEN, virtual histology evaluation of atherosclerosis regression during atorvastatin and ezetimibe administration study; PRECISE-IVUS, Plaque Regression With Cholesterol Absorption Inhibitor or Synthesis Inhibitor Evaluated by Intravascular Ultrasound Study; IMPROVE-IT, the Improved Reduction of Outcomes: Vytorin Efficacy International Trial.

statistically significant difference between the two pooled RRs (in the diabetes and no diabetes groups).

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When all included trials were analysed by meta-regression, there was a trend toward a greater MACE risk reduction by ezetimibe combination therapy when added to statins in patients with diabetes compared to those without diabetes (β =0.89; 95% CI, 0.75 to 1.06; *P*=0.203). When the placebo-controlled study was excluded from the analysis, ezetimibe-statin combination therapy was associated with a greater reduction of MACE risk in subjects with diabetes than in those without diabetes compared with statin monotherapy (β =0.87; 95% CI, 0.78 to 0.99; *P*=0.038).

Data on cancer incidence were available in three studies (Kouvelos, SHARP, and IMPROVE-IT) (Supplemental Fig. S2). The pooled RR for cancer incidence was 1.01 (95% CI,

 Table 2. Sensitivity Analyses for Assessing the Effects of Individual Studies on the Pooled Risk Ratio for Major Adverse Cardiovascular Events

	Pooled RR	95% CI
Omitted study (DM) ^a		
West	0.84	0.77-0.91
SHARP	0.86	0.78-0.94
HEAVEN	0.84	0.77-0.91
Suzuki	0.84	0.77-0.91
Kouvelos	0.85	0.78-0.93
PRECISE-IVUS	0.84	0.77-0.91
IMPROVE-IT	0.77	0.65-0.92
Combined	0.84	0.77-0.91
Omitted studies (non-DM) ^b		
West	0.91	0.80-1.02
SHARP	0.97	0.91-1.03
HEAVEN	0.91	0.81-1.02
Suzuki	0.92	0.83-1.02
Kouvelos	0.96	0.91-1.01
PRECISE-IVUS	0.94	0.86-1.02
IMPROVE-IT	0.85	0.74-0.97
Combined	0.93	0.85-1.02

RR, risk ratio; CI, confidence interval; DM, diabetes mellitus; SHARP, the Study of Heart and Renal Protection; HEAVEN, Virtual histology evaluation of atherosclerosis regression during atorvastatin and ezetimibe administration study; PRECISE-IVUS, Plaque Regression With Cholesterol Absorption Inhibitor or Synthesis Inhibitor Evaluated by Intravascular Ultrasound Study; IMPROVE-IT, The Improved Reduction of Outcomes: Vytorin Efficacy International Trial.

^aDM, subgroup with diabetes in each study; ^bNon-DM, subgroup without diabetes in each study.

0.94 to 1.09; P=0.794), indicating no difference between the ezetimibe and control groups.

Sensitivity analysis

In the sensitivity analysis (Table 2), the meta-analyses were repeated after removing one study at a time. Omitting individual trials did not significantly affect the pooled risk, and the risk reduction by ezetimibe combination therapy in the diabetes group remained robust even after removal of the largest trial (subanalvsis excluding the IMPROVE-IT trial: [RR, 0.77; 95% CI, 0.65 to 0.92 vs. RR, 0.84; 95% CI, 0.77 to 0.91; P=0.372 in the diabetes subgroup]; [RR, 0.85; 95% CI, 0.74 to 0.97 vs. RR, 0.93; 95% CI, 0.85 to 1.02; P=0.278 in the non-diabetes subgroups]). However, the difference between the pooled RRs in the diabetes and non-diabetes groups did not reach statistical significance when the IMPROVE-IT trial was excluded (RR, 0.78; 95% CI, 0.65 to 0.93 in the diabetes subgroup vs. RR, 0.85; 95% CI, 0.74 to 0.97 in the non-diabetes group; $P_{\text{heterogeneity}} = 0.460$), although a nominally consistent pattern was observed with the pooled results of the studies overall.

Publication bias

A funnel plot and Egger test of the studies did not reveal any evidence of underlying publication bias for reporting MACEs (Supplemental Fig. S3).

DISCUSSION

In this meta-analysis of seven RCTs, we found a differential association of ezetimibe-statin combination therapy on MACE risk according to the presence of diabetes. Compared with statins alone, ezetimibe combination therapy reduced the risk of MACEs. The benefit of ezetimibe combination therapy was more prominent in patients with diabetes than in patients without diabetes.

Recent reviews and meta-analyses of RCTs, including the IMPROVE-IT study, showed that ezetimibe was likely associated with a reduction of the risk of myocardial infarction and stroke, without affecting the risk of overall or cardiovascular mortality or newly-developed cancer [21,22]. However, published reviews reported marginal cardiovascular benefits of ezetimibe when ezetimibe was added to statins for reducing nonfatal myocardial infarction and stroke (17 fewer myocardial infarctions and six fewer strokes per 1,000 persons treated over 6 years) [21-24]. This uncertainty is reflected in the absence of a consensus regarding ezetimibe in international guidelines. The

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2013 treatment guidelines of the American Heart Association/ American College of Cardiology (AHA/ACC) focused on statin monotherapy and did not suggest considering second-line drugs, including ezetimibe, as a treatment option based on a lack of strong evidence [4]. However, European and Korean guidelines for the management of dyslipidaemia permit the use of ezetimibe as a second-line therapy in association with statins when the therapeutic goal is not met despite maximal tolerated statin doses or in subjects intolerant to statins [25,26]. In line with newer evidence, the 2016 AHA/ACC updates on cholesterol treatment commented that second-line cholesterol-lowering drugs can be used to meet LDL-C treatment targets, at least in limited circumstances [27]. Given the current evidence of the ability of ezetimibe to prevent cardiovascular events, it is important to identify the specific populations that might benefit the most from ezetimibe. However, no reviews or meta-analyses have primarily focused on the differential effect of ezetimibe according to the presence of diabetes. In this study, the pooled results of RCTs with a statin control arm showed that patients with diabetes experienced a greater benefit from ezetimibestatin combination therapy than patients without diabetes, indicating that the presence of diabetes might be a potential indication for adding ezetimibe to the therapeutic regimen of patients with high residual risk.

Several biological and clinical findings support the beneficial effect of ezetimibe in diabetes. Patients with a higher risk of cardiovascular events experience greater benefits when ezetimibe is added to statins, as shown in a previous review [22]. Patients with diabetes are more likely to have a higher cardiovascular risk at baseline than patients without diabetes, which might lead to ezetimibe exerting a positive effect in patients with diabetes [8]. Furthermore, pathologic enhancement of NPC1L1 expression, a direct target of ezetimibe, has been reported in patients with diabetes [28,29]. Indeed, ezetimibe was associated with greater decreases in LDL-C and non-high density cholesterol levels in patients with diabetes than in those without diabetes [30,31]. In addition to its favourable effects on the lipid profile of individuals with diabetes, ezetimibe combination therapy was associated with improvements in insulin sensitivity and plasma adiponectin levels compared with statin monotherapy in patients with diabetes [32]. In the Plaque Regression With Cholesterol Absorption Inhibitor or Synthesis Inhibitor Evaluated by Intravascular Ultrasound Study (PRE-CISE-IVUS) trial, which was included in this analysis, the greater reduction of atherosclerotic plaque progression by ezetimibe could not be entirely explained by its cholesterol-lowering effects [20]. Anti-inflammatory effects, reduction of the plant sterol ratio, inhibition of smooth muscle cell proliferation, and antiplatelet effects have been proposed as potential mechanisms underlying the cardiovascular benefit of ezetimibe [33-36]. Taken together, these pieces of evidence suggest that ezetimibe combination therapy might have a protective effect on cardiovascular outcomes in patients with diabetes, possibly through its lipid-lowering effects or through a pleiotropic effect; however, the mechanisms of the effects of ezetimibe must be confirmed in further studies.

Concerns regarding increased cancer-related mortality associated with ezetimibe use have been raised. According to a pooled analysis of three large trials (Simvastatin in Aortic Stenosis [SEAS], SHARP, and IMPROVE-IT), ezetimibe use was associated with a nominally increased risk of cancer-related mortality (risk ratio, 1.45; 99% CI, 1.02 to 2.05; uncorrected P=0.007) [37,38]. However, the authors argued that this result might have been due to chance, rather than being a true finding, because a parallel increase in the cancer incidence was not found in the combined analysis. In our study, we also observed no association of ezetimibe use with cancer incidence when SHARP, IM-PROVE-IT, and the study by Kouvelos et al. [18] were pooled together, similarly to the meta-analysis performed by Savarese et al. [21]. Although monitoring for mortality due to cancer should be continued in large prospective trials, our findings support the current consensus that ezetimibe is most likely not associated with an increased risk of cancer incidence.

Our study is limited by the small number of eligible RCTs, with a single study representing the majority of enrolled patients. Although we intentionally only analysed RCTs to minimize heterogeneity, it is possible that excluding observational studies with large numbers of subjects and longer follow-up durations might have led to an underestimation of the effect size of ezetimibe. Surrogate outcomes were not analysed in this study. Although we analysed LDL-C levels according to the treatment groups, we could not obtain changes in the LDL-C level for each study stratified by diabetes. A composite endpoint, MAC-Es, was analysed instead of individual outcomes due to the lack of data stratified by the presence of diabetes, although the scope of this study was to evaluate the heterogeneity of the effects of ezetimibe on cardiovascular events between individuals with diabetes and those without diabetes. The studies that remained after excluding the SHARP and the IMPROVE-IT trials in our study might have been underpowered for detecting a significant additive cardioprotective effect of ezetimibe combination therapy between the diabetes and non-diabetes groups. Meanwhile, a

previous meta-analysis of the effects of ezetimibe emphasized that including large studies such as SHARP or IMPROVE-IT in the pooled outcome analyses led to a significantly larger sample of patients, with greater representativeness of real-world patients, when compared to a meta-analysis performed without the results from large trials [21,23]. Furthermore, our study provided a comparison of the pooled risk of cardiovascular outcomes between diabetes and no diabetes groups, based on data that were collected by direct contact. Therefore, we believe that this analysis contributes some novel information on the interaction of the effects of ezetimibe with the presence of diabetes, although further prospective trials are needed to validate this possibility.

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In conclusion, the pooled results of RCTs showed that ezetimibe was associated with a greater reduction of MACEs in patients with diabetes than in those without diabetes. This differential effect of ezetimibe was robust across the trials. Given the current evidence regarding ezetimibe as a second-line lipidlowering agent, ezetimibe-statin combination therapy might provide a feasible treatment option to combat residual cardiovascular risk in patients with diabetes who are intolerant or refractory to statin therapy.

CONFLICTS OF INTEREST

No potential conflict of interest relevant to this article was reported.

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AUTHOR CONTRIBUTIONS

All authors had full access to all of the data in the study and can take responsibility for the integrity of the data and the accuracy of the data analysis. Designed the study and wrote the protocol: N.H., Y.H.L., E.S.K. Supervised data collection and synthesis: E.S.K. Wrote the report and final draft of the manuscript: Y.H.L., N.H. Wrote the search strategy and undertook the literature search: K.H. Undertook all data analysis: Y.H.L., N.H. Undertook title screening, data gathering, cleaning and advised on methods, statistical analyses, and interpretation of the findings: K.T., J.A.G., C.M.K., T.K., G.N.K., H.S., C.J.L., S.H.P., B. W.L., B.S.C. Contributed equally to this work: N.H., Y.H.L. Guarantor: E.S.K. All authors contributed to the final manuscript.

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